SECTION 4: CURVE FITTING

ESC 440 – Computational Methods for Engineers



Curve Fitting

Often, we have data, y, that is a function of some independent variable, x,

Possibly noisy measurement data

- Underlying relationship is unknown
 - Know x's and y's (approximately)
 - But, don't know y = f(x)



Curve Fitting

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- May want to determine a function (i.e., a curve) that 'best' describes relationship between x and y
 - An approximation to (the unknown) y = f(x)
 - This is *curve fitting*



Regression vs. Interpolation

We'll look at two categories of curve fitting:

Least-squares regression

Noisy data – uncertainty in y value for a given x value
 Want "good" agreement between f(x) and data points
 Curve (i.e., f(x)) may not pass through any data points

Polynomial interpolation

Data points are known exactly – noiseless data
 Resulting curve passes through all data points

6 Review of Basic Statistics

Before moving on to discuss least-squares regression, we'll first review a few basic concepts from statistics.

Basic Statistical Quantities

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- □ *Arithmetic mean* the average or expected value

$$\overline{y} = \frac{\sum y_i}{n}$$

 Standard deviation (unbiased) – a measure of the spread of the data about the mean

$$\sigma = \sqrt{\frac{S_t}{n-1}}$$

where S_t is the **total sum of the squares of the residuals**

$$S_t = \sum (y_i - \bar{y})^2$$

Basic Statistical Quantities

- □ *Variance* another measure of spread
 - The square of the standard deviation
 - Useful measure due to relationship with power and power spectral density of a signal or data set

$$\sigma^2 = \frac{S_t}{n-1} = \frac{\sum (y_i - \bar{y})^2}{n-1}$$

or

$$\sigma^2 = \frac{\sum y_i^2 - \frac{(\sum y_i)^2}{n}}{n-1}$$

Normal (Gaussian) Distribution

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- Many naturally-occurring random process are normally-distributed
 - Measurement noise
 - Very often assume noise in our data is Gaussian
 - Probability density function (pdf):

$$f(x) = \frac{1}{\sqrt{2\pi\sigma^2}} e^{-\frac{(x-\mu)^2}{2\sigma^2}}$$

where σ^2 is the variance, and μ is the mean of the random variable, x

- Very often useful to generate *random numbers* Simulating the effect of noise
 Monte Carlo simulation, etc.
- First, construct a random-number generator object using NumPy:

rng = np.random.default_rng(seed)

 seed: *optional* initialization seed for generator
 rng: initialized generator object – will run methods on this object to generate random numbers

Normally-Distributed Random Numbers

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- Generate random values from a normal (Gaussian) distribution
 - x = rng.normal(loc=0, scale=1, size=1)
 - rng: generator object created with default_rng()
 loc: *optional* mean of distribution default: 0.0
 scale: *optional* standard deviation default: 1.0
 size: *optional* dimension of resulting array
 x: resulting array of random values

Uniformly-Distributed Random Numbers

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- Generate random values from a uniform distribution on the interval [low, high)

x = rng.uniform(low=0, high=1, size=1)

- ng: generator object created with default_rng()
- Iow: optional lower bound of interval default: 0.0
- high: optional upper bound of interval default: 1.0
- size: optional dimension of resulting array default: 1
- x: resulting array of random values
- □ Half-open interval:
 - Resulting values are ≥ low and < high</p>

NumPy Statistical Functions

- NumPy includes many statistical functions, including:
 - np.max()
 - □ np.min()
 - np.mean()
 - np.std()
 - np.median()
 - np.var()
 - □ np.cov()

Histogram Plots

Histogram plots

Graphical depiction of the variation of random quantities

- Plots the frequency of occurrence of ranges (bins) of values
- Provides insight into the nature of the distribution

plt.hist(x, bins=20, edgecolor='k')

- **•** x: data to be histogrammed
- bins: optional number of bins
- edgecolor: optional color of bin outlines default: none

Statistics in NumPy, matplotlib







¹⁶ Linear Least-Squares Regression

Linear Regression

- Noisy data, y, values at known x values
- Suspect relationship
 between x and y is
 linear
- i.e., assume
 - $y = a_0 + a_1 x$
- Determine a₀ and a₁
 that define the "*best-fit*" line for the data



Assumed a linear relationship between x and y:

$$y = a_0 + a_1 x$$

Due to noise, can't measure y exactly at each x Can only approximate y values

$$\hat{y} = y + e$$

Measured values are approximations
 True value of y plus some random error or residual

$$\hat{y} = a_0 + a_1 x + e$$

Best Fit Criteria

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- Noisy data do not all line on a single line discrepancy between each point and the line fit to the data
 The error, or *residual*:

$$e = \hat{y} - a_0 - a_1 x$$

- Minimize some measure of this residual:
 - Minimize the sum of the residuals
 - Positive and negative errors can cancel
 - Non-unique fit
 - Minimize the sum of the absolute values of the residuals
 - Effect of sign of error eliminated, but still not a unique fit
 - Minimize the maximum error minimax criterion
 - Excessive influence given to single outlying points

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 Better fitting criterion is to minimize the *sum of the* squares of the residuals

$$S_r = \sum e_i^2 = \sum (\hat{y}_i - a_0 - a_1 x_i)^2$$

Yields a unique best-fit line for a given set of data

- □ The sum of the squares of the residuals is a function of the two fitting parameters, a_0 and a_1 , $S_r(a_0, a_1)$
- Minimize S_r by setting its partial derivatives to zero and solving for a₀ and a₁

Least-Squares Criterion

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- □ At its minimum point, partial derivatives of S_r with respect to a_0 and a_1 will be zero

$$\frac{\partial S_r}{\partial a_0} = -2\sum (\hat{y}_i - a_0 - a_1 x_i) = 0$$
$$\frac{\partial S_r}{\partial a_1} = -2\sum [(\hat{y}_i - a_0 - a_1 x_i) x_i] = 0$$

□ Breaking up the summation:

$$\sum \hat{y}_i - \sum a_0 - \sum a_1 x_i = 0$$
$$\sum x_i \hat{y}_i - \sum a_0 x_i - \sum a_1 x_i^2 = 0$$

Normal Equations

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□ $\partial S_r / \partial a_0 = 0$ and $\partial S_r / \partial a_1 = 0$ form a system of two equations with two unknowns, a_0 and a_1

$$n a_0 + \left(\sum x_i\right) a_1 = \sum \hat{y}_i \tag{1}$$

$$\left(\sum x_i\right)a_0 + \left(\sum x_i^2\right)a_1 = \sum x_i\hat{y}_i \tag{2}$$

In matrix form:

$$\begin{bmatrix} n & \sum x_i \\ \sum x_i & \sum x_i^2 \end{bmatrix} \begin{bmatrix} a_0 \\ a_1 \end{bmatrix} = \begin{bmatrix} \sum \hat{y}_i \\ \sum x_i \hat{y}_i \end{bmatrix}$$
(3)

These are the *normal equations*

Normal Equations

Normal equations can be solved for a_0 and a_1 :

$$a_1 = \frac{n \sum x_i \hat{y}_i - \sum x_i \sum \hat{y}_i}{n \sum x_i^2 - (\sum x_i)^2}$$
$$a_0 = \frac{\sum \hat{y}_i - a_1 \sum x_i}{n} = \bar{y} - a_1 \bar{x}$$

 Or solve the matrix form of the normal equations, (3), in Python using np.linalg.solve()

Linear Least-Squares - Example

- 24
- Noisy data with
 suspected linear
 relationship
- Calculate summation terms in the normal equations:

$$\square$$
 n, Σx_i , $\Sigma \hat{y}_i$, Σx_i^2 , $\Sigma x_i \hat{y}_i$

22	n = len(yn)
23	Sx = sum(x)
24	Sy = sum(yn)
25	Sxy = sum(x*yn)
26	$Sx2 = sum(x^{**}2)$



Linear Least-Squares - Example

- Assemble normal equation matrices
- Solve normal equations for vector of coefficients, a, using np.linalg.solve()

22	<pre>n = len(yn)</pre>
23	Sx = sum(x)
24	Sy = sum(yn)
25	Sxy = sum(x*yn)
26	$Sx2 = sum(x^{**}2)$
27	
28	<pre>Z = np.array([[n, Sx], [Sx, Sx2]])</pre>
29	<pre>b = np.array([Sy, Sxy])</pre>
30	<pre>a = np.linalg.solve(Z, b)</pre>
31	
32	# %% the best-fit line
33	y1 = a[1]*x + a[0]
7.4	



Goodness of Fit

- How well does a function fit the data?
- Is a linear fit best? A quadratic, higher-order polynomial, or other non-linear function?
- Want a way to be able to quantify goodness of fit
- Quantify spread of data about the mean prior to regression:

$$S_t = \sum (\hat{y}_i - \bar{y})^2$$

Following regression, quantify spread of data about the regression line (or curve):

$$S_r = \sum (\hat{y}_i - a_0 - a_1 x_i)^2$$

Goodness of Fit

- $\Box S_t$ quantifies the spread of the data about the mean
- *S_r* quantifies spread about the best-fit line (curve)
 The spread that remains after the trend is explained
 The *unexplained sum of the squares*
- $\Box S_t S_r$ represents the reduction in data spread after regression explains the underlying trend
- Normalize to S_t the *coefficient of determination*

$$r^2 = \frac{S_t - S_r}{S_t}$$

Coefficient of Determination

$$r^2 = \frac{S_t - S_r}{S_t}$$

□ For a perfect fit:

No variation in data about the regression line

$$\Box S_r = 0 \quad \rightarrow \quad r^2 = 1$$

If the fit provides no improvement over simply characterizing data by its mean value:

$$\Box S_r = S_t \quad \rightarrow \quad r^2 = 0$$

If the fit is worse at explaining the data than their mean value:

 $\square S_r > S_t \quad \rightarrow \quad r^2 < 0$

Coefficient of Determination

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Calculate r² for previous example:

39	<pre># calculate the coefficient</pre>
40	# of determination
41	ybar = np.mean(yn)
42	<pre>St = sum((yn - ybar)**2)</pre>
43	$Sr = sum((yn - y1)^{**2})$
44	r2 = (St - Sr)/St
ALC:	





Coefficient of Determination

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- Don't rely too heavily on the value of r²
 Anscombe's famous data sets:



Same line fit to all four data sets
 r² = 0.67 in each case

³¹ Linearization of Nonlinear Relationships

Nonlinear functions

 Not all data can be explained by a linear relationship to an independent variable, e.g.

Exponential model

$$y = \alpha e^{\beta x}$$

Power equation

$$y = \alpha x^{\beta}$$

Saturation-growth-rate equation

$$y = \alpha \frac{x}{\beta + x}$$

Nonlinear functions

Methods for nonlinear curve fitting:

Linearization of the nonlinear relationship

- Transform the dependent and/or independent data values
- Apply linear least-squares regression
- Inverse transform the determined coefficients back to those that define the nonlinear functional relationship

Nonlinear regression

■ Treat as an optimization problem – more later...

Linearizing an Exponential Relationship

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 Have noisy data that is believed to be best described by an *exponential relationship*

Dac

$$y = \alpha e^{\beta x}$$

$$Linearize the fittingequation:
$$\ln(y) = \ln(\alpha) + \beta x$$
or

$$\ln(y) = a_0 + a_1 x$$
where

$$a_0 = \ln(\alpha), a_1 = \beta$$
Measured Data
Measured Data$$

Linearizing an Exponential Relationship

- 35
- Fit a line to the transformed data using linear leastsquares regression
- □ Determine a_0 and a_1 :
 - $\ln(y) = a_0 + a_1 x$
- Can calculate r² for the line fit to the transformed data
- Note that original data must be positive



Linearizing an Exponential Relationship

- 36
- Transform the linear fitting parameters, a₀ and a₁, back to the parameters defining the exponential relationship

0

□ Exponential fit:



the line fit to the

transformed data



Best-Fit Exponential
Linearizing an Exponential Relationship



51	# %% inverse transform the linear
52	# fit coefficients to get the
53	# parameters for the exponential
54	alpha = np.exp(a[0])
55	beta = a[1]
56	
57	# the exponential fit
58	<pre>yexp = alpha*np.exp(beta*xfit)</pre>
59	
60	<pre># calculate the coefficient</pre>
61	<pre># of determination for exp fit</pre>
62	ybar = np.mean(yn)
63	$St = sum((yn - ybar)^{**2})$
64	Sr = sum((yn - yexp[-len(x):])**2)
65	r2 = (St - Sr)/St



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Have noisy data that is believed to be best described by an *power equation*

$$y = \alpha x^{\beta}$$
Linearize the fitting
equation:
$$\log(y) = \log(\alpha) + \beta \log(x) ,$$
or
$$\log(y) = a_0 + a_1 \log(x)$$

where

$$a_0 = \log(\alpha), \ a_1 = \beta$$



- 39
- Fit a line to the transformed data using linear leastsquares regression
- \Box Determine a_0 and a_1 :

 $\log(y) = a_0 + a_1 \log(x)$

- Can calculate r² for the line fit to the transformed data
- Note that original data both x and y – must be positive



- 40
- Transform the linear fitting parameters, a₀ and a₁, back to the parameters defining the power equation
- Power equation:

 $y = \alpha x^{\beta}$

where

$$lpha=10^{a_0}$$
 , $\ \ eta=a_1$

 Note that r² is different than that for the line fit to the transformed data





51	# %% inverse transform the linear
52	<pre># fit coefficients to get the</pre>
53	<pre># parameters for the power equation</pre>
54	alpha = 10**(a[0])
55	beta = a[1]
56	
57	# the power equation fit
58	<pre>xpow = np.arange(max(x)+1)</pre>
59	ypow = alpha*xpow**beta
60	<pre>ypowr2 = alpha*x**beta</pre>
61	
62	<pre># calculate the coefficient of</pre>
63	<pre># determination for power eqn. fit</pre>
64	ybar = np.mean(yn)
65	$St = sum((yn - ybar)^{**2})$
66	$Sr = sum((yn - ypowr2)^{**2})$
67	r2 = (St - Sr)/St





Have noisy data that is believed to be best described by a saturation growth-rate equation

$$y = \alpha \frac{x}{\beta + x}$$

□ Linearize the fitting equation:

$$\frac{1}{y} = \frac{1}{\alpha} + \frac{\beta}{\alpha} \frac{1}{x}$$

or

$$\frac{1}{y} = a_0 + a_1 \frac{1}{x}$$

where

$$a_0 = \frac{1}{\alpha}$$
, $a_1 = \frac{\beta}{\alpha}$



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- Fit a line to the transformed data using linear leastsquares regression
- \Box Determine a_0 and a_1 :

$$\frac{1}{y} = a_0 + a_1 \frac{1}{x}$$

 Can calculate r² for the line fit to the transformed data



- 44
- Transform the linear fitting parameters, a₀ and a₁, back to the parameters defining the saturation growth-rate equation
- Saturation growth-rate equation:

$$y = \alpha \frac{x}{\beta + x}$$

where

$$\alpha = \frac{1}{a_0}, \quad \beta = \frac{a_1}{a_0}$$

 Note that r² is different than that for the line fit to the transformed data







Best-Fit Saturation Growth-Rate Equation



46 Polynomial Regression

Polynomial Regression

- So far we've looked at fitting straight lines to *linear* and *linearized* data sets
- Can also fit *mth-order polynomials* directly to data using *polynomial regression*
- Same fitting criterion as linear regression:
 - Minimize the sum of the squares of the residuals
 - m+1 fitting parameters for an mth-order polynomial
 - m+1 normal equations

Polynomial Regression

- Assume, for example, that we have data we believe to be *quadratic* in nature
- □ 2nd-order polynomial regression
- □ Fitting equation:

$$\hat{y} = a_0 + a_1 x + a_2 x^2 + e$$

Best fit will minimize the sum of the squares of the residuals:

$$S_r = \sum (\hat{y}_i - a_0 - a_1 x_i - a_2 x_i^2)^2$$

Polynomial Regression – Normal Equations

Best-fit polynomial coefficients will minimize S_r
 Differentiate S_r w.r.t. each coefficient and set to zero

$$\frac{\partial S_r}{\partial a_0} = -2\sum \left(\hat{y}_i - a_0 - a_1 x_i - a_2 x_i^2\right) = 0$$

$$\frac{\partial S_r}{\partial a_1} = -2\sum x_i (\hat{y}_i - a_0 - a_1 x_i - a_2 x_i^2) = 0$$

 $\frac{\partial S_r}{\partial a_2} = -2\sum x_i^2 (\hat{y}_i - a_0 - a_1 x_i - a_2 x_i^2) = 0$

Polynomial Regression – Normal Equations

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Rearranging the normal equations yields

$$n a_{0} + (\Sigma x_{i})a_{1} + (\Sigma x_{i}^{2})a_{2} = \Sigma \hat{y}_{i}$$

$$(\Sigma x_{i})a_{0} + (\Sigma x_{i}^{2})a_{1} + (\Sigma x_{i}^{3})a_{2} = \Sigma x_{i}\hat{y}_{i}$$

$$(\Sigma x_{i}^{2})a_{0} + (\Sigma x_{i}^{3})a_{1} + (\Sigma x_{i}^{4})a_{2} = \Sigma x_{i}^{2}\hat{y}_{i}$$

Which can be put into matrix form:

$$\begin{bmatrix} n & \Sigma x_i & \Sigma x_i^2 \\ \Sigma x_i & \Sigma x_i^2 & \Sigma x_i^3 \\ \Sigma x_i^2 & \Sigma x_i^3 & \Sigma x_i^4 \end{bmatrix} \begin{bmatrix} a_0 \\ a_1 \\ a_2 \end{bmatrix} = \begin{bmatrix} \Sigma \hat{y}_i \\ \Sigma x_i \hat{y}_i \\ \Sigma x_i^2 \hat{y}_i \end{bmatrix}$$

This system of equations can be solved for the vector of unknown coefficients using NumPy's linalg.solve()

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- For *mth-order* polynomial regression the *normal* equations are:

$$\begin{bmatrix} n & \Sigma x_i & \cdots & \Sigma x_i^m \\ \Sigma x_i & \Sigma x_i^2 & \cdots & \Sigma x_i^{m+1} \\ \vdots & \vdots & \ddots & \vdots \\ \Sigma x_i^m & \Sigma x_i^{m+1} & \cdots & \Sigma x_i^{2m} \end{bmatrix} \begin{bmatrix} a_0 \\ a_1 \\ \vdots \\ a_m \end{bmatrix} = \begin{bmatrix} \Sigma \hat{y}_i \\ \Sigma x_i \hat{y}_i \\ \vdots \\ \Sigma x_i^m \hat{y}_i \end{bmatrix}$$

Again, this system of m + 1 equations can be solved for the vector of m + 1 unknown polynomial coefficients using NumPy's linalg.solve()

Polynomial Regression – Example

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6	# %% noiseless data
7	<pre>p = np.poly([1,3,9])</pre>
8	<pre>x = np.linspace(0,10,25)</pre>
9	<pre>y = np.polyval(p,x)</pre>
10	
11	# %% add noise to y data
12	sig = 8
13	
14	# set the random number generator seed
15	seed = 4
16	
17	<pre>rng = np.random.default_rng(seed)</pre>
18	<pre>v = rng.normal(scale=sig, size=len(x))</pre>
19	yn = y + v

Best-Fit Cubic



21	# %% Construct and solve the
22	# normal equations
23	n = len(yn)
24	Sx = sum(x)
25	$Sx2 = sum(x^{**2})$
26	$Sx3 = sum(x^{**}3)$
27	$5x4 = sum(x^{**}4)$
28	$Sx5 = sum(x^{**5})$
29	$Sx6 = sum(x^{**}6)$
30	Sy = sum(yn)
31	Sxy = sum(x*yn)
32	$Sx2y = sum(x^{**}2^*yn)$
33	$Sx3y = sum(x^{**}3^{*}yn)$
34	
35	Z = [[n, Sx, Sx2, Sx3],
36	[Sx, Sx2, Sx3, Sx4],
37	[Sx2, Sx3, Sx4, Sx5],
38	[Sx3, Sx4, Sx5, Sx6]]
39	b = [Sy, Sxy, Sx2y, Sx3y]
40	<pre>a = np.linalg.solve(Z,b)</pre>
41	
42	# %% reverse the order of coefficients to
43	<pre># conform to NumPy's convention</pre>
44	<pre>pfit = a[::-1]</pre>
45	# or
46	<pre>pfit = np.flip(a)</pre>
47	
48	<pre># np.polyfit will give the same answer</pre>
49	<pre># pfit = np.polyfit(x,yn,3)</pre>
50	
51	# %% evaluate the best-fit cubic
52	<pre>xfit = np.linspace(min(x),max(x),200)</pre>
53	<pre>y3 = np.polyval(pfit,xfit)</pre>
54	y3r2 = np.polyval(pfit,x)
55	
56	# %% calculate the coefficient
57	<pre># of determination for the fit</pre>
58	ybar = np.mean(yn)
59	$St = sum((yn - ybar)^{**2})$
60	$Sr = sum((yn - y3r2)^{**2})$
61	r2 = (St - Sr)/St

Polynomial Regression - np.polyfit()

x: n-vector of independent variable data values
 y: n-vector of dependent variable data values
 m: order of the polynomial to be fit to the data
 p: (m + 1)-vector of best-fit polynomial coefficients

- Least-squares polynomial regression if:
 - $\square n > m + 1$
 - **I** i.e., for over-determined systems
- Polynomial interpolation if:
 - $\square n = m + 1$
 - Resulting fit passes through all (x,y) points more later

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Polynomial Regression - np.polyfit()



 Note that the result matches that obtained by solving normal equations

```
# polyfit3.py
 З
       import numpy as np
 4
       from matplotlib import pyplot as plt
 5
 6
 7
       # %% noiseless data
 8
       p = np.poly([1,3,9])
       x = np.linspace(0, 10, 25)
 9
       y = np.polyval(p,x)
10
11
12
       # %% add noise to y data
13
       sig = 8
14
15
       # set the random number generator seed
16
       seed = 4
17
18
       rng = np.random.default rng(seed)
       v = rng.normal(scale=sig, size=len(x))
19
20
       yn = y + v
21
       # %% use np.polyfit to perform the regression
22
       pfit = np.polyfit(x,yn,3)
23
24
       # %% evaluate the best-fit cubic
25
       xfit = np.linspace(min(x),max(x),200)
26
       y3 = np.polyval(pfit,xfit)
27
28
       y3r2 = np.polyval(pfit,x)
29
30
       # %% calculate the coefficient
       # of determination for the fit
31
       ybar = np.mean(yn)
32
33
       St = sum((yn - ybar)^{**2})
34
       Sr = sum((yn - y3r2)^{**2})
       r2 = (St - Sr)/St
35
```

Polynomial Regression Using np.polyfit()

- Determine the 4th-order polynomial with roots at $x = \{1, 5, 16, 19\}$
- Generate noiseless data points by evaluating this polynomial at integer values of x from 0 to 20
- Add Gaussian white noise with a standard deviation of $\sigma = 180$ to your data points
- Use np.polyfit() to fit a 4th-order polynomial to the noisy data
- Calculate the coefficient of determination, r^2
- Plot the noisy data points, along with the best-fit polynomial



Multiple Linear Regression

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- We have so far fit lines or curves to data described by functions of a single variable
- For functions of multiple variables, *fit planes or surfaces to data*
- Linear function of two independent variables: *multiple linear regression*

$$\hat{y} = a_0 + a_1 x_1 + a_2 x_2 + e$$

Sum of the squares of the residuals is now

$$S_r = \sum \left(\hat{y}_i - a_0 - a_1 x_{1,i} - a_2 x_{2,i} \right)^2$$

Multiple Linear Regression – Normal Equations

- Differentiate S_r w.r.t. fitting coefficients and equate to zero
- □ The *normal equations*:

$$\begin{bmatrix} n & \Sigma x_{1,i} & \Sigma x_{2,i} \\ \Sigma x_{1,i} & \Sigma x_{1,i}^2 & \Sigma x_{1,i} x_{2,i} \\ \Sigma x_{2,i} & \Sigma x_{1,i} x_{2,i} & \Sigma x_{2,i}^2 \end{bmatrix} \begin{bmatrix} a_0 \\ a_1 \\ a_2 \end{bmatrix} = \begin{bmatrix} \Sigma \hat{y}_i \\ \Sigma x_{1,i} \hat{y}_i \\ \Sigma x_{2,i} \hat{y}_i \end{bmatrix}$$

Solve as before – now fitting coefficients, a_i, define a *plane*



- We've seen three types of least-squares regression
 Linear regression
 Polynomial regression
 Multiple linear regression
- All are special cases of *general linear least-squares regression*

$$\hat{y} = a_0 z_0 + a_1 z_1 + \dots + a_m z_m + e$$

- \Box The z_i 's are m + 1 **basis functions**
 - Basis functions may be nonlinear
 - This is *linear* regression, because dependence on fitting coefficients, a_i, is linear

$$\hat{y} = a_0 z_0 + a_1 z_1 + \dots + a_m z_m + e$$

□ For *linear regression* – simple or multiple: $z_0 = 1, z_1 = x_1, z_2 = x_2, ... z_m = x_m$

□ For *polynomial regression*:

$$z_0 = 1, z_1 = x, z_2 = x^2, \dots z_m = x^m$$

In all cases, this is a *linear combination of basis function*, which may, themselves, be *nonlinear*

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The general linear least-squares model:

$$\hat{y} = a_0 z_0 + a_1 z_1 + \dots + a_m z_m + e$$

□ Can be expressed in matrix form:

$$\hat{\mathbf{y}} = \mathbf{Z} \mathbf{a} + \mathbf{e}$$

where **Z** is an $n \times (m + 1)$ matrix, the **design matrix**, whose entries are the (m + 1) basis functions evaluated at the n independent variable values corresponding to the n measurements:

$$\mathbf{Z} = \begin{bmatrix} z_{01} & z_{11} & \cdots & z_{m1} \\ z_{02} & z_{12} & \cdots & z_{m2} \\ \vdots & \vdots & \ddots & \vdots \\ z_{0n} & z_{1n} & \cdots & z_{mn} \end{bmatrix}$$

where z_{ij} is the i^{th} basis function evaluated at the j^{th} independent variable value. (Note: *i* is not the row index and *j* is not the column index, here.)

The least-squares model is:

$$\begin{bmatrix} z_{01} & z_{11} & \cdots & z_{m1} \\ z_{02} & z_{12} & \cdots & z_{m2} \\ \vdots & \vdots & \ddots & \vdots \\ z_{0n} & z_{1n} & \cdots & z_{mn} \end{bmatrix} \begin{bmatrix} a_0 \\ a_1 \\ \vdots \\ a_m \end{bmatrix} + \begin{bmatrix} e_1 \\ e_2 \\ \vdots \\ e_n \end{bmatrix} = \begin{bmatrix} \hat{y}_1 \\ \hat{y}_2 \\ \vdots \\ \hat{y}_n \end{bmatrix}$$

More measurements than coefficients

- $\square n > (m+1)$
- **Z** is not square tall and narrow
- Over-determined system
- Z⁻¹ does not exist

General Linear Least-Squares – Design Matrix Example

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□ For example, consider fitting a quadratic to five measured values, $\hat{\mathbf{y}}$, at $\mathbf{x} = [1, 2, 3, 4, 5]^T$

Model is:

$$\hat{y} = a_0 + a_1 x + a_2 x^2 + e$$

Basis functions are $z_0 = 1$, $z_1 = x$, and $z_2 = x^2$ Least-squares equation is

$$\begin{bmatrix} 1 & 1 & 1 \\ 1 & 2 & 4 \\ 1 & 3 & 9 \\ 1 & 4 & 16 \\ 1 & 5 & 25 \end{bmatrix} \begin{bmatrix} a_0 \\ a_1 \\ a_2 \end{bmatrix} = \begin{bmatrix} \hat{y}_1 \\ \hat{y}_2 \\ \hat{y}_3 \\ \hat{y}_4 \\ \hat{y}_5 \end{bmatrix} - \begin{bmatrix} e_1 \\ e_2 \\ e_3 \\ e_4 \\ e_5 \end{bmatrix}$$

General Linear Least-Squares – Residuals

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Linear least-squares model is:

$$\hat{\mathbf{y}} = \mathbf{Z} \, \mathbf{a} + \mathbf{e} \tag{1}$$

Residual:

$$\mathbf{e} = \hat{\mathbf{y}} - \mathbf{y} = \hat{\mathbf{y}} - \mathbf{Z} \mathbf{a}$$
(2)

Sum of the squares or the residuals:

$$S_{r} = \sum e_{i}^{2} = \mathbf{e}^{\mathsf{T}}\mathbf{e} = [\hat{\mathbf{y}} - \mathbf{Z} \mathbf{a}]^{\mathsf{T}}[\hat{\mathbf{y}} - \mathbf{Z} \mathbf{a}] \qquad (3)$$

Expanding,

$$S_{r} = \hat{\mathbf{y}}^{\mathsf{T}} \hat{\mathbf{y}} - \mathbf{a}^{\mathsf{T}} \mathbf{Z}^{\mathsf{T}} \hat{\mathbf{y}} - \hat{\mathbf{y}}^{\mathsf{T}} \mathbf{Z} \mathbf{a} + \mathbf{a}^{\mathsf{T}} \mathbf{Z}^{\mathsf{T}} \mathbf{Z} \mathbf{a}$$
(4)

Deriving the Normal Equations

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- Best fit will minimize the sum of the squares of the residuals
 - Differentiate S_r with respect to the coefficient vector, a, and set to zero

$$\frac{dS_{r}}{d\mathbf{a}} = \frac{d}{d\mathbf{a}} \left(\hat{\mathbf{y}}^{\mathsf{T}} \hat{\mathbf{y}} - \mathbf{a}^{\mathsf{T}} \mathbf{Z}^{\mathsf{T}} \hat{\mathbf{y}} - \hat{\mathbf{y}}^{\mathsf{T}} \mathbf{Z} \mathbf{a} + \mathbf{a}^{\mathsf{T}} \mathbf{Z}^{\mathsf{T}} \mathbf{Z} \mathbf{a} \right) = \mathbf{0} \quad (5)$$

We'll need to use some *matrix calculus identities:*

$$\Box \frac{d}{da} (a^{T} Z^{T} y) = Z^{T} y$$
$$\Box \frac{d}{da} (y^{T} Z a) = Z^{T} y$$
(6)
$$\Box \frac{d}{da} (a^{T} Z^{T} Z a) = 2Z^{T} Z a$$

Deriving the Normal Equations

$$\frac{dS_{r}}{d\mathbf{a}} = \frac{d}{d\mathbf{a}} \left(\hat{\mathbf{y}}^{\mathsf{T}} \hat{\mathbf{y}} - \mathbf{a}^{\mathsf{T}} \mathbf{Z}^{\mathsf{T}} \hat{\mathbf{y}} - \hat{\mathbf{y}}^{\mathsf{T}} \mathbf{Z} \mathbf{a} + \mathbf{a}^{\mathsf{T}} \mathbf{Z}^{\mathsf{T}} \mathbf{Z} \mathbf{a} \right) = \mathbf{0}$$

Using the matrix derivative relationships, (6),

$$\frac{dS_r}{d\mathbf{a}} = -2\mathbf{Z}^{\mathsf{T}}\hat{\mathbf{y}} + 2\mathbf{Z}^{\mathsf{T}}\mathbf{Z}\mathbf{a} = \mathbf{0}$$
(7)

Equation (7) is the matrix form of the *normal* equations:

$$\mathbf{Z}^{\mathsf{T}}\mathbf{Z}\mathbf{a} = \mathbf{Z}^{\mathsf{T}}\hat{\mathbf{y}}$$
(8)

Solution to (8) is the vector of least-squares fitting coefficients:

$$\mathbf{a} = \left(\mathbf{Z}^{\mathrm{T}}\mathbf{Z}\right)^{-1}\mathbf{Z}^{\mathrm{T}}\widehat{\mathbf{y}}$$
(9)

Solving the Normal Equations

$$\mathbf{a} = \left(\mathbf{Z}^{\mathsf{T}}\mathbf{Z}\right)^{-1}\mathbf{Z}^{\mathsf{T}}\hat{\mathbf{y}}$$
(9)

Remember, our starting point was the linear leastsquares model:

$$\mathbf{y} = \mathbf{Z} \mathbf{a} \tag{10}$$

Couldn't we have solved (10) for fitting coefficients as

$$\mathbf{a} = \mathbf{Z}^{-1}\mathbf{y} \tag{11}$$

- □ No, must solve using (9), because:
 - $\hfill\square$ Don't have y, only noisy approximations, \widehat{y}
 - We have an over-determined system
 - **Z** is not square
 - Z⁻¹ does not exist

Solving the Normal Equations

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Solution to the linear least-squares problem is:

$$\mathbf{a} = \left(\mathbf{Z}^{\mathsf{T}}\mathbf{Z}\right)^{-1}\mathbf{Z}^{\mathsf{T}}\hat{\mathbf{y}} = \mathbf{Z}^{\dagger}\hat{\mathbf{y}}$$
(12)

where

$$\mathbf{Z}^{\dagger} = \left(\mathbf{Z}^{\mathsf{T}}\mathbf{Z}\right)^{-1}\mathbf{Z}^{\mathsf{T}}$$
(13)

is the *Moore-Penrose pseudo-inverse* of Z

Use the pseudo-inverse to find the least-squares solutions to an over-determined system

Coefficient of Determination

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Goodness of fit characterized by the *coefficient of determination:*

$$r^2 = \frac{S_t - S_r}{S_t}$$

where S_r is given by (3)

$$S_{r} = [\hat{\mathbf{y}} - \mathbf{Z} \, \mathbf{a}]^{\mathrm{T}} [\hat{\mathbf{y}} - \mathbf{Z} \, \mathbf{a}]$$
(14)

and

$$S_t = [\hat{\mathbf{y}} - \bar{\mathbf{y}}]^{\mathsf{T}} [\hat{\mathbf{y}} - \bar{\mathbf{y}}]$$
(15)

General Least-Squares in Python

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Have *n* measurements

$$\hat{y} = [\hat{y}_0 \quad \hat{y}_1 \quad \cdots \quad \hat{y}_{n-1}]^T$$

 \Box at *n* known independent variable values

$$x = \begin{bmatrix} x_0 & x_1 & \cdots & x_{n-1} \end{bmatrix}^T$$

 \Box and a model, defined by m + 1 basis functions

$$\hat{y} = a_0 z_0 + a_1 z_1 + \dots + a_m z_m + e$$

Generate design matrix by evaluating m + 1 basis functions at all n values of x

$$\mathbf{Z} = \begin{bmatrix} z_0(x_0) & z_1(x_0) & \cdots & z_m(x_0) \\ z_0(x_1) & z_1(x_1) & \cdots & z_m(x_1) \\ \vdots & \vdots & \ddots & \vdots \\ z_0(x_{n-1}) & z_1(x_{n-1}) & \cdots & z_m(x_{n-1}) \end{bmatrix}$$

General Least-Squares in Python

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Solve for vector of fitting coefficients as the solution to the normal equations

$$\mathbf{a} = \left(\mathbf{Z}^{\mathsf{T}}\mathbf{Z}\right)^{-1}\mathbf{Z}^{\mathsf{T}}\hat{\mathbf{y}}$$

□ Or by using np.linalg.lstsq()

a = np.linalg.lstsq(Z, yhat)

Result is the same, though the methods are different
73 Nonlinear Regression

Nonlinear Regression - minimize()

Nonlinear models:

Have nonlinear dependence on fitting parameters
E.g., $y = \alpha x^{\beta}$

- Two options for fitting nonlinear models to data
 - Linearize the model first, then use linear regression
 - Fit a nonlinear model directly by treating as an optimization problem
- Want to *minimize* a *cost function* Cost function is the *sum of the squares of the residuals*

$$J = S_r = \sum (\hat{y} - y)^2$$

Find the minimum of J – a *multi-dimensional optimization* Use SciPy's optimize.minimize()

Nonlinear Regression - minimize()

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- Have noisy data that is believed to be best described by an *exponential relationship*

$$y = \alpha e^{\beta x}$$

Cost function:

$$J = \sum \left(\hat{y} - \alpha e^{\beta x}\right)^2$$

 Find α and β to minimize J
 Use SciPy's optimize.minimize()



Find the minimum of a function of two or more variables

- **f**: function to be optimized
- x0: array of initial values
- opt: optimizeResult object returned includes:
 - opt.x: the solution of the optimization (i.e., x_{opt})
 - opt.fun: value of objective function at the optimum (i.e., $f(x_{opt})$)
 - opt.nit: number of iterations

Nonlinear Regression - minimize()

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Nonlinear Regression - curve_fit()

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- An alternative to minimizing a cost function using scipy.optimize.curve_fit():

 f: handle to the fitting function – independent variable must be listed first

- **•** x: independent variable data
- y: dependent variable data
- p0: initial guess for popt optional
- popt: best-fit parameters
- pcov: estimated covariance of popt

Nonlinear Regression - curve_fit()

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⁸⁰ Polynomial Interpolation

Polynomial Interpolation

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- Sometimes we know both x and y values exactly
 Want a function that describes y = f(x)
 Allows for *interpolation* between know data points
 - **\square** Fit an n^{th} -order polynomial to n + 1 data points

$$y = a_0 x^n + a_1 x^{n-1} + a_2 x^{n-2} + \dots + a_n$$

Polynomial will pass through all points

- We'll look at *polynomial interpolation* using
 - Newton's polynomial
 - The *Lagrange polynomial*

Polynomial Interpolation

$$y = a_0 x^n + a_1 x^{n-1} + a_2 x^{n-2} + \dots + a_n$$

□ Can approach similar to linear least-squares regression $y = a_0 z_0 + a_1 z_1 + \dots + a_n z_n$

where

$$z_0 = x^n$$
, $z_1 = x^{n-1}$, ... $z_n = 1$

■ For an n^{th} -order polynomial, we have n + 1 equations with n + 1 unknowns

In matrix form

$$\mathbf{y} = \mathbf{Z} \mathbf{a}$$

Polynomial Interpolation

- Now, unlike for linear regression
 - All n + 1 values in **y** are known exactly
 - n + 1 equations with n + 1 unknown coefficients
 - **Z** is square $(n + 1) \times (n + 1)$

$$\begin{bmatrix} y_1 \\ y_2 \\ \vdots \\ y_{n+1} \end{bmatrix} = \begin{bmatrix} x_1^n & x_1^{n-1} & \cdots & 1 \\ x_2^n & x_2^{n-1} & \cdots & 1 \\ \vdots & \vdots & \ddots & \vdots \\ x_{n+1}^n & x_{n+1}^{n-1} & \cdots & 1 \end{bmatrix} \begin{bmatrix} a_0 \\ a_1 \\ \vdots \\ a_n \end{bmatrix}$$

Could solve by inverting Z or by using NumPy's linalg.solve()

Z is a Vandermonde matrix

- Tend to be ill-conditioned
- The techniques that follow are more numerically robust

⁸⁴ Newton Interpolating Polynomial

Linear Interpolation

Fit a line (1st-order polynomial) to two data points using a truncated *Taylor series* (or simple trigonometry):

$$f_1(x) = f(x_1) + \frac{f(x_2) - f(x_1)}{x_2 - x_1}(x - x_1)$$

where $f_1(x)$ is the function for the line fit to the data, and $f(x_i)$ are the known data values

This is the Newton linear-interpolation formula

Quadratic Interpolation

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To fit a 2nd-order polynomial to three data points, consider the following form

$$f_2(x) = b_0 + b_1(x - x_1) + b_2(x - x_1)(x - x_2)$$

 \square Evaluate at $x = x_1$ to find b_0

$$b_0 = f(x_1)$$

Back-substitution and evaluation at $x = x_2$ and at $x = x_3$ will yield the other coefficients

$$b_1 = \frac{f(x_2) - f(x_1)}{x_2 - x_1}$$
 and $b_2 = \frac{\frac{f(x_3) - f(x_2)}{x_3 - x_2} \frac{f(x_2) - f(x_1)}{x_2 - x_1}}{x_3 - x_1}$

Quadratic Interpolation

$$f_2(x) = b_0 + b_1(x - x_1) + b_2(x - x_1)(x - x_2)$$

- Can still view this as a Taylor series approximation
 - **\square** b_0 represents an offset
 - $\blacksquare b_1$ is slope
 - \square b_2 is curvature
- Choice of initial quadratic form (Newton interpolating polynomial) was made to facilitate the development
 - Resulting polynomial would be the same for any initial form of an nth-order polynomial
 - Solution is unique

*n*th-Order Newton Interpolating Polynomial

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Extending the quadratic example to nth-order

$$f_n(x) = b_0 + b_1(x - x_1) + \dots + b_n(x - x_1)(x - x_2) \dots (x - x_n)$$

□ Solve for coefficients as before with back-substitution and evaluation of $f(x_i)$

$$b_{0} = f(x_{1})$$

$$b_{1} = f[x_{2}, x_{1}]$$

$$b_{2} = f[x_{3}, x_{2}, x_{1}]$$

$$\vdots$$

$$b_{n} = f[x_{n+1}, x_{n}, \dots, x_{2}, x_{1}]$$

$\Box f[\cdots]$ denotes a *finite divided difference*

Finite Divided Differences

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First finite divided difference

$$f[x_i, x_j] = \frac{f(x_i) - f(x_j)}{x_i - x_j}$$

Second finite divided difference

$$f[x_i, x_j, x_k] = \frac{f[x_i, x_j] - f[x_j, x_k]}{x_i - x_k}$$

□ n^{th} finite divided difference $f[x_{n+1}, x_n, ..., x_2, x_1] = \frac{f[x_{n+1}, ..., x_2] - f[x_n, ..., x_1]}{f[x_n, ..., x_1]}$

$$x_{n+1} - x_1$$

Calculate recursively

K. Webb

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- nth-order Newton interpolating polynomial in terms of divided differences:

$$f_n(x) = f(x_1) + f[x_2, x_1](x - x_1) + \cdots$$

+ $f[x_{n+1}, x_n, \dots, x_2, x_1](x - x_1)(x - x_2) \cdots (x - x_n)$

Divided difference table for calculation of coefficients:



Newton Interpolating Polynomial – Example

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def newtpoly(x,y,xint):

```
Calculates an interpolating polynomial through
points in x and y, using a Newton polynomial.
Order of the polynomial is n=length(x)-1
Parameters
x : array of independent variable data
y : array of dependent variable data
xint : array of independent variable values at which interpolation is
       performed
Returns
yint: array of values of the interpolating polynomial evaluated at xint
# initialize the finite divided difference matrix
                # order of polynomial is n-1
n = len(x)
fdd = np.zeros((n,n))
fdd[:,0] = y
# recursively calculate all divided differences
for j in range(1, n): # column index steps from col 1 to col n-1
    for i in range(0, n-j): # row index steps from 1 down to the off diag.
        fdd[i,j] = (fdd[i+1,j-1] - fdd[i,j-1])/(x[i+j] - x[i])
# first row of fdd are the b_i coefficients
\# b = fdd(0, -1; -1; -1 - n)
b = fdd[0,:]
# generate interpolating polynomial
yint = np.zeros(len(xint))
                        # first term in the polynomial sum at x=xint[0]
yint[0] = b[0]
# loop through all values of xint, interpolating at each
for k in range(len(xint)):
    xt = 1
    yint[k] = b[0]
    for m in range(1, n):
        # xt is product of (x-xi) terms - one more for each incr. of m
        xt = (xint[k] - x[m-1])*xt
       yint[k] = yint[k] + b[m]*xt
return yint
```



Linear Lagrange Interpolation

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Fit a first-order polynomial (a line) to two known data points: $(x_1, f(x_1))$ and $(x_2, f(x_2))$

$$f_1(x) = L_1(x) \cdot f(x_1) + L_2(x) \cdot f(x_2)$$

 \Box $L_1(x)$ and $L_2(x)$ are **weighting functions**, where

$$L_{1}(x) = \begin{cases} 1, & x = x_{1} \\ 0, & x = x_{2} \end{cases}$$
$$L_{2}(x) = \begin{cases} 1, & x = x_{2} \\ 0, & x = x_{1} \end{cases}$$

The interpolating polynomial is a weighted sum of the individual data point values

Linear Lagrange Interpolation

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For linear (1st-order) interpolation, the weighting functions are:

$$L_1(x) = \frac{x - x_2}{x_1 - x_2}$$
$$L_2(x) = \frac{x - x_1}{x_2 - x_1}$$

The linear Lagrange interpolating polynomial is:

$$f_1(x) = \frac{x - x_2}{x_1 - x_2} f(x_1) + \frac{x - x_1}{x_2 - x_1} f(x_2)$$

nth-Order Lagrange Interpolation

 Lagrange interpolation technique can be extended to nth-order polynomials

$$f_n(x) = \sum_{i=1}^{n+1} L_i(x) \cdot f(x_i)$$

where

$$L_{i}(x) = \prod_{\substack{j=1 \ j \neq i}}^{n+1} \frac{x - x_{j}}{x_{i} - x_{j}}$$

Lagrange Interpolating Polynomial – Example

1	# lagpoly_test.py
2	
3	import numpy as np
4	<pre>from matplotlib import pyplot as plt</pre>
5	<pre>from poly_interp import lagpoly</pre>
6	
7	x = [1, 2, 3, 4]
8	y = [20,26,15,-2]
9	
10	<pre>xint = np.linspace(x[0],x[-1],500)</pre>
11	
12	<pre>yint = lagpoly(x,y,xint)</pre>
13	



