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Math 8652: Theory of Probability

Homework #5

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① Let $X_1, \dots, X_n \in L^1(\Omega, \mathcal{F}, \mathbb{P})$. Denote by $\text{Span}(X_1, \dots, X_n)$ the linear span of $\{X_1, X_2, \dots, X_n\}$ in $L^1(\Omega, \mathcal{F}, \mathbb{P})$. We show by induction in $n \in \mathbb{N}$ that $\text{Span}\{X_1, \dots, X_n\}$ is closed in $L^1(\Omega, \mathcal{F}, \mathbb{P})$.

For $n=1$, $\text{Span}(X_1) = \{cX_1 : c \in \mathbb{R}\}$. If $X_1 = 0$ a.s. then $\text{Span}(X_1) = \{0\}$, which is closed in $L^1(\Omega, \mathcal{F}, \mathbb{P})$. Consider the case X_1 is not almost surely equal to 0. For each $m \in \mathbb{N}$, put

$$A_m = \left\{ \omega \in \Omega : X_1(\omega) > \frac{1}{m} \right\},$$

$$B_m = \left\{ \omega \in \Omega : X_1(\omega) < -\frac{1}{m} \right\}.$$

Then $(\bigcup_{m=1}^{\infty} A_m) \cup (\bigcup_{m=1}^{\infty} B_m) = \{\omega \in \Omega : X_1(\omega) \neq 0\}$, which has positive measure

Then $0 < \mathbb{P}\left(\left(\bigcup_{m=1}^{\infty} A_m\right) \cup \left(\bigcup_{m=1}^{\infty} B_m\right)\right) \leq \sum_{m=1}^{\infty} \mathbb{P}(A_m) + \sum_{m=1}^{\infty} \mathbb{P}(B_m)$.

This implies there exists $m_0 \in \mathbb{N}$ such that $\mathbb{P}(A_{m_0}) > 0$ or $\mathbb{P}(B_{m_0}) > 0$.

By replacing X_1 with $-X_1$ if necessary, we can assume $\mathbb{P}(A_{m_0}) > 0$.

Let (Z_m) be a sequence in $\text{Span}(X_1)$ that converges to $Z \in L^1(\Omega, \mathcal{F}, \mathbb{P})$.

Write $Z_m = c_m X_1$. Then

$$\int_{\Omega} Z_m \mathbb{I}_{A_{m_0}} \mathbb{P}(d\omega) = c_m \underbrace{\int_{\Omega} X_1 \mathbb{I}_{A_{m_0}} \mathbb{P}(d\omega)}_a. \quad (1)$$

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$$a = \int_{\Omega} X_1 I_{A_{m_0}} \mathbb{P}(d\omega) \geq \frac{1}{m_0} \mathbb{P}(A_{m_0}) > 0.$$

$$\left| \int_{\Omega} (Z_m - Z) I_{A_{m_0}} \mathbb{P}(d\omega) \right| \leq \int_{\Omega} |Z_m - Z| \mathbb{P}(d\omega) = \|Z_m - Z\|_{L^1} \rightarrow 0 \text{ as } m \rightarrow \infty.$$

Thus,

$$\lim_{m \rightarrow \infty} \int_{\Omega} Z_m I_{A_{m_0}} \mathbb{P}(d\omega) = \int_{\Omega} Z I_{A_{m_0}} \mathbb{P}(d\omega).$$

Then (1) implies

$$c_m = \frac{1}{a} \int_{\Omega} Z_m I_{A_{m_0}} \mathbb{P}(d\omega) \longrightarrow \frac{1}{a} \int_{\Omega} Z I_{A_{m_0}} \mathbb{P}(d\omega) := c \text{ as } m \rightarrow \infty.$$

Then $Z_m = c_m X_1 \rightarrow c X_1$ pointwise as $m \rightarrow \infty$. For all m sufficiently large, $|c_m - c| \leq 1$. Then $|Z_m - c X_1| = |c_m - c| |X_1| \leq |X_1|$, which is an integrable function. By the Dominated Convergence Theorem, $Z_m \rightarrow c X_1$ in $L^1(\Omega, \mathcal{F}, \mathbb{P})$.

Thus, $Z = c X_1 \in \text{Span}(X_1)$.

Let $n \geq 2$. Suppose $\text{Span}(X_1, \dots, X_r)$ is closed in $L^1(\Omega, \mathcal{F}, \mathbb{P})$ for any $1 \leq r < n$ and $X_1, X_2, \dots, X_r \in L^1(\Omega, \mathcal{F}, \mathbb{P})$. Take $X_1, X_2, \dots, X_n \in L^1(\Omega, \mathcal{F}, \mathbb{P})$. We show that $\text{Span}(X_1, \dots, X_n)$ is closed in $L^1(\Omega, \mathcal{F}, \mathbb{P})$. Let (\tilde{X}_m) be a sequence in $\text{Span}(X_1, \dots, X_n)$ that converges to some $X \in L^1(\Omega, \mathcal{F}, \mathbb{P})$.

Because $\text{Span}(X_1, \dots, X_n)$ is a finite dimensional vector space, it has a basis $\{Y_1, Y_2, \dots, Y_s\}$ for some $0 \leq s \leq n$. If $s = 0$ then $\text{Span}(X_1, \dots, X_n) = \{0\}$, which is closed in $L^1(\Omega, \mathcal{F}, \mathbb{P})$. If $s = 1$ then $\text{Span}(X_1, \dots, X_n) = \text{Span}(Y_1)$, which has been proved to be closed in $L^1(\Omega, \mathcal{F}, \mathbb{P})$. Consider the case $s \geq 2$.

Write

$$\tilde{X}_m = \alpha_{1m} Y_1 + \alpha_{2m} Y_2 + \dots + \alpha_{sm} Y_s$$

for $\alpha_{1m}, \alpha_{2m}, \dots, \alpha_{sm} \in \mathbb{R}$. For each $j \in \{1, 2, \dots, s\}$, we show that the sequence $(\alpha_{jm})_{m \in \mathbb{N}}$ is bounded. Suppose otherwise. Without loss of generality, we can assume the sequence $(\alpha_{sm})_{m \in \mathbb{N}}$ is unbounded. By replacing (\tilde{X}_m) with a suitable subsequence, we can assume $|\alpha_{sm}| \rightarrow \infty$ as $m \rightarrow \infty$. Then

$$\begin{aligned} Y_s &= \frac{1}{\alpha_{sm}} (\tilde{X}_m - \alpha_{1m} Y_1 - \dots - \alpha_{s-1,m} Y_{s-1}) \\ &= \frac{1}{\alpha_{sm}} \tilde{X}_m + \sum_{j=1}^{s-1} \left(-\frac{\alpha_{jm}}{\alpha_{sm}} \right) Y_j \quad \forall m \in \mathbb{N}. \quad (2) \end{aligned}$$

Because $\tilde{X}_m \rightarrow X$ in $L^1(\Omega, \mathcal{F}, \mathbb{P})$ and $|\alpha_{sm}| \rightarrow \infty$ as $m \rightarrow \infty$, $\frac{1}{\alpha_{sm}} \tilde{X}_m \rightarrow 0$ in $L^1(\Omega, \mathcal{F}, \mathbb{P})$. Letting $m \rightarrow \infty$ in (2), we get

$$Y_s = \lim_{m \rightarrow \infty} \sum_{j=1}^{s-1} \left(-\frac{\alpha_{jm}}{\alpha_{sm}} \right) Y_j.$$

This implies Y_s belongs to the closure in $L^1(\Omega, \mathcal{F}, \mathbb{P})$ of $\text{Span}(Y_1, \dots, Y_{s-1})$.

Since $1 \leq s-1 < n$, by the induction hypothesis, $\text{Span}(Y_1, \dots, Y_{s-1})$ is closed in $L^1(\Omega, \mathcal{F}, \mathbb{P})$. Then $Y_s \in \text{Span}(Y_1, \dots, Y_{s-1})$. This is a contradiction because Y_1, Y_2, \dots, Y_s are linearly independent.

We have showed that the sequence $(\alpha_{jm})_{m \in \mathbb{N}}$ is bounded for every $j \in \{1, 2, \dots, s\}$. (α_{1m}) has a convergent subsequence (α_{1m_k}) . Denote by $\tilde{X}^{(1)}$ the subsequence (\tilde{X}_{m_k}) of (\tilde{X}_m) . If we regard $\tilde{X}^{(1)}$ instead of (\tilde{X}_m) , we can assume (α_{1m}) converges. (α_{2m}) has a convergent subsequence $(\alpha_{2m'_k})$. Denote by

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$\tilde{X}^{(2)}$ the subsequence $(\tilde{X}_{m_k}^{(1)})$ of $\tilde{X}^{(1)}$. If we regard $\tilde{X}^{(2)}$ instead of (\tilde{X}_m) , we can assume (α_{1m}) and (α_{2m}) converge. Continue this process s times. We then conclude that (\tilde{X}_m) has a subsequence such that $(\alpha_{1m}), (\alpha_{2m}), \dots, (\alpha_{sm})$ converge as $m \rightarrow \infty$. Replace (\tilde{X}_m) by this subsequence. Put $\alpha_j = \lim_{m \rightarrow \infty} \alpha_{jm}$ for each $j \in \{1, 2, \dots, s\}$. Put $Y = \alpha_1 Y_1 + \dots + \alpha_s Y_s$. Then

$$\tilde{X}_m = \alpha_{1m} Y_1 + \dots + \alpha_{sm} Y_s \longrightarrow \alpha_1 Y_1 + \dots + \alpha_s Y_s = Y \text{ pointwise as } m \rightarrow \infty.$$

$$|\tilde{X}_m - Y| \leq \sum_{j=1}^s |\alpha_{jm} - \alpha_j| |Y_j| \leq \sum_{j=1}^s |Y_j| \in L^1(\Omega, \mathcal{F}, \mathbb{P})$$

for all m sufficiently large. By the Dominated Convergence Theorem, $\tilde{X}_m \rightarrow Y$ in $L^1(\Omega, \mathcal{F}, \mathbb{P})$. Hence, $X = Y \in \text{Span}(Y_1, \dots, Y_s) = \text{Span}(X_1, \dots, X_n)$.

② Let $\Omega = \bigcup_{n=1}^{\infty} A_n$ be a partition of Ω into disjoint sets $A_n \in \mathcal{F}$ such that $\mathbb{P}(A_n) > 0$ for every $n \in \mathbb{N}$. Put $\mathcal{G} = \sigma(A_n : n \in \mathbb{N})$. Take $\xi \in L^1(\Omega, \mathcal{F}, \mathbb{P})$. We show that

$$E(\xi | \mathcal{G}) = \frac{1}{\mathbb{P}(A_n)} E(\xi I_{A_n})$$

on A_n for every $n \in \mathbb{N}$. Equivalently, we need to show

$$E(\xi | \mathcal{G}) = \sum_{n=1}^{\infty} \frac{1}{\mathbb{P}(A_n)} E(\xi I_{A_n}) I_{A_n}. \quad (1)$$

Denote

$$X = \sum_{n=1}^{\infty} \frac{1}{\mathbb{P}(A_n)} E(\xi I_{A_n}) I_{A_n},$$

$$X_m = \sum_{n=1}^m \frac{1}{\mathbb{P}(A_n)} E(\xi I_{A_n}) I_{A_n}.$$

The indicators $I_{A_1}, I_{A_2}, I_{A_3}, \dots$ are \mathcal{G} -measurable. Thus, X_m is \mathcal{G} -measurable.

Because A_1, A_2, A_3, \dots are pairwise disjoint, $X_m = X$ on $\bigcup_{n=1}^m A_n$. ~~Thus,~~
~~Because~~ $\bigcup_{n=1}^{\infty} A_n = \Omega$,

$$X(\omega) = \lim_{n \rightarrow \infty} X_n(\omega) \quad \forall \omega \in \Omega.$$

Then X is also \mathcal{G} -measurable.

Next, we show that $X \in L^1(\Omega, \mathcal{F}, \mathbb{P})$ and $EX = E\beta$.

$$\begin{aligned} \int_{\Omega} |X_m| \mathbb{P}(d\omega) &\leq \int_{\Omega} \sum_{n=1}^m \frac{1}{\mathbb{P}(A_n)} E(\beta | I_{A_n}) I_{A_n} \mathbb{P}(d\omega) \\ &= \sum_{n=1}^m \frac{1}{\mathbb{P}(A_n)} E(\beta | I_{A_n}) \mathbb{P}(A_n) \\ &= \sum_{n=1}^m E(\beta | I_{A_n}) \\ &= E(\beta | I_{\bigcup_{n=1}^m A_n}). \end{aligned} \tag{2}$$

Then
$$\begin{aligned} \int_{\Omega} |X| \mathbb{P}(d\omega) &\stackrel{\text{Fatou}}{\leq} \liminf_{m \rightarrow \infty} \int_{\Omega} |X_m| \mathbb{P}(d\omega) \leq \limsup_{m \rightarrow \infty} E(\beta | I_{\bigcup_{n=1}^m A_n}) \\ &= E(\beta | I_{\Omega}) \quad (\text{by Monotone Convergence Theorem}) \\ &= E\beta < \infty. \end{aligned}$$

While we was deriving (2), we also derived $|X_m| \leq \beta |I_{\bigcup_{n=1}^m A_n}| \leq \beta$.

Thus, by the Dominated Convergence Theorem,

$$\begin{aligned} EX &= \int_{\Omega} X \mathbb{P}(d\omega) = \lim_{m \rightarrow \infty} \int_{\Omega} X_m \mathbb{P}(d\omega) = \lim_{m \rightarrow \infty} \int_{\Omega} \sum_{n=1}^m \frac{1}{\mathbb{P}(A_n)} E(\beta | I_{A_n}) I_{A_n} \\ &= \lim_{m \rightarrow \infty} \sum_{n=1}^m \frac{1}{\mathbb{P}(A_n)} E(\beta | I_{A_n}) \mathbb{P}(A_n) \end{aligned}$$

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$$\begin{aligned}
 &= \lim_{m \rightarrow \infty} \sum_{n=1}^m E(\xi I_{A_n}) = \lim_{m \rightarrow \infty} E(\xi I_{\bigcup_{n=1}^m A_n}) \stackrel{11A}{=} \\
 &= E(\xi I_{\Omega}) \quad (\text{by the Dominated Convergence Theorem}) \\
 &= E\xi.
 \end{aligned}$$

To show that $X = E(\xi | \mathcal{G})$, we need to show

$$E(X I_A) = E(\xi I_A) \quad \forall A \in \mathcal{G}.$$

Put $\mathcal{G}_0 = \{A \in \mathcal{G} : E(X I_A) = E(\xi I_A)\}$. We now show that \mathcal{G}_0 is a π -system.

Since $E(X I_\phi) = E(\xi I_\phi) = 0$, $\phi \in \mathcal{G}_0$. Let B_1, B_2, B_3, \dots be a sequence of pairwise disjoint members of \mathcal{G}_0 .

$$E(X I_{B_n}) = E(\xi I_{B_n}) \quad \forall n \in \mathbb{N}.$$

Then

$$E(X I_{\bigcup_{n=1}^m B_n}) = \sum_{n=1}^m E(X I_{B_n}) = \sum_{n=1}^m E(\xi I_{B_n}) = E(\xi I_{\bigcup_{n=1}^m B_n}) \quad \forall m \in \mathbb{N}. \quad (3)$$

Put $B = \bigcup_{n=1}^{\infty} B_n \in \mathcal{G}$. Then $|X I_{\bigcup_{n=1}^m B_n}| \leq |X|$. By the Dominated Convergence Theorem,

$$\lim_{m \rightarrow \infty} E(X I_{\bigcup_{n=1}^m B_n}) = E(X I_B).$$

$$\text{Similarly, } \lim_{m \rightarrow \infty} E(\xi I_{\bigcup_{n=1}^m B_n}) = E(\xi I_B).$$

Letting $m \rightarrow \infty$ in (3), we get $E(X I_B) = E(\xi I_B)$. Thus $B \in \mathcal{G}_0$.

Take any $A \in \mathcal{G}_0$. Then

$$E(X I_{A^c}) = E(X - X I_A) = EX - E(X I_A),$$

$$E(\xi I_{A^c}) = E(\xi - \xi I_A) = E\xi - E(\xi I_A).$$

Because $EX = E\xi$ and $E(XI_A) = E(\xi I_A)$, we get $E(XI_{A^c}) = E(\xi I_{A^c})$.
Then $A^c \in \mathcal{G}_0$. We have showed that \mathcal{G}_0 is a λ -system.

For each $k \in \mathbb{N}$, $XI_{A_k} = \frac{1}{P(A_k)} E(\xi I_{A_k}) I_{A_k}$ because

$$I_{A_k} I_{A_n} = I_{A_k \cap A_n} = \begin{cases} I_{A_k} & \text{if } n=k, \\ 0 & \text{if } n \neq k. \end{cases}$$

$$\begin{aligned} \text{Thus, } E(XI_{A_k}) &= \int_{\Omega} \frac{1}{P(A_k)} E(\xi I_{A_k}) I_{A_k} P(d\omega) = \frac{1}{P(A_k)} E(\xi I_{A_k}) P(A_k) \\ &= E(\xi I_{A_k}). \end{aligned}$$

This means $A_k \in \mathcal{G}_0$. The family $\alpha = \{\emptyset, A_1, A_2, A_3, \dots\}$ is a π -system because A_1, A_2, A_3, \dots are pairwise disjoint. Then α is a π -system contained in the λ -system \mathcal{G}_0 . By the Sierpinski Class theorem, \mathcal{G}_0 contains $\sigma(\alpha) = \sigma(A_1, A_2, A_3, \dots) = \mathcal{G}$. Thus, $\mathcal{G}_0 = \mathcal{G}$.

③ Let $d \in \mathbb{N}$ and $a_{ij}, b_i, c \in \mathbb{R}$ for $1 \leq i, j \leq d$. Suppose the function $f: \mathbb{R}^d \rightarrow \mathbb{R}$

$$f(x) = \sum_{i,j=1}^d a_{ij} x_i x_j + 2 \sum_{i=1}^d b_i x_i + c$$

is nonnegative for every $x \in \mathbb{R}^d$. We show that f has a minimum value.

Denote by $a = (a_{ij})_{1 \leq i, j \leq d} \in M_d(\mathbb{R})$, $b = (b_i)_{1 \leq i \leq d} \in \mathbb{R}^d$. Denote by (\cdot, \cdot)

the usual dot product in \mathbb{R}^d . Then

$$f(x) = (ax, x) + 2(b, x) + c. \quad (1)$$

We observe that $\sum_{i,j=1}^d a_{ij} x_i x_j = \sum_{i,j=1}^d \frac{a_{ij} + a_{ji}}{2} x_i x_j$.

This means if we replace matrix a with $\frac{a+a^T}{2}$, the value of $f(x)$ does

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not change. Thus, we can assume a is symmetric, i.e. $a_{ji} = a_{ij}$ for all $1 \leq i, j \leq d$.

$$f(tx) = t^2(a x, x) + 2t(b, x) + c \geq 0 \quad \forall t \in \mathbb{R}. \quad (2)$$

If $(a x, x) < 0$ for some $x \in \mathbb{R}^d$ then (2) is not satisfied as $t \rightarrow \infty$. Thus, $(a x, x) \geq 0$ for all $x \in \mathbb{R}^d$. We view LHS(2) as a quadratic polynomial in t .

The discriminant must be nonpositive, i.e.

$$(b, x)^2 \leq c(a x, x) \quad \forall x \in \mathbb{R}^d.$$

Because a is symmetric, the transpose matrix of a is $a^* = a$. For each $x \in \ker a^*$,

$$(b, x)^2 \leq c(a x, x) = c(a^* x, x) = c(0, x) = 0.$$

Thus, $(b, x) = 0$. Then $b \in (\ker a^*)^\perp$. In linear algebra, we know the identity

$$\text{range}(a) = (\ker a^*)^\perp.$$

Thus, $b \in \text{range}(a)$. There exists $x_0 \in \mathbb{R}^d$ such that $b = a x_0$. Then (1) gives

$$\begin{aligned} f(x) &= (a x, x) + 2(a x_0, x) + c \\ &= (a x, x) + (a x_0, x) + (a x_0, x) + c \\ &= (a x, x) + (a x_0, x) + \underbrace{(x_0, a x)}_{=(a x, x_0)} + c \\ &= (a x + a x_0, x + x_0) - (a x_0, x_0) + c \\ &= \underbrace{(a(x+x_0), x+x_0)}_{\geq 0} - (a x_0, x_0) + c \\ &\geq -(a x_0, x_0) + c. \end{aligned}$$

The equality holds when $x = -x_0$. Therefore, $\min_{x \in \mathbb{R}^d} f(x) = f(-x_0)$.

④ Consider a random walk with steps X_1, X_2, X_3, \dots . Suppose $P(X_1=1) = P(X_1=-1) = \frac{1}{2}$. Put $S_n = X_1 + X_2 + \dots + X_n$, and $\xi_n = \left(\frac{5}{4}\right)^{-n} 2^{S_n}$. First, we show that $\xi_n \rightarrow 0$ a.s.

We have $EX_1 = 1P(X_1=1) + (-1)P(X_1=-1) = \frac{1}{2} - \frac{1}{2} = 0$. By the Strong law of large numbers, $\frac{S_n}{n} \rightarrow EX_1 = 0$ a.s. Then

$$\log \xi_n = S_n \log 2 - n \log \frac{5}{4} = n \log 2 \left(\underbrace{\frac{S_n}{n} - \frac{\log \frac{5}{4}}{\log 2}}_{\rightarrow -\frac{\log \frac{5}{4}}{\log 2} < 0} \right) \rightarrow -\infty \text{ a.s.}$$

Thus, $\xi_n = e^{\log \xi_n} \rightarrow 0$ a.s.

Next, we show $E \sup_{n \in \mathcal{N}} \xi_n = \infty$. Put $\xi = \sup_{n \in \mathcal{N}} \xi_n \geq 0$. Suppose by contradiction that $E\xi < \infty$. Because $0 \leq \xi_n \leq \xi$ and $\xi_n \rightarrow 0$ a.s., by the Dominated Convergence Theorem,

$$\lim_{n \rightarrow \infty} E\xi_n = E(\lim_{n \rightarrow \infty} \xi_n) = E(0) = 0. \tag{1}$$

On the other hand,

$$\begin{aligned} E\xi_n &= \left(\frac{5}{4}\right)^{-n} 2^{S_n} = \left(\frac{5}{4}\right)^{-n} E(2^{X_1} 2^{X_2} \dots 2^{X_n}) \\ &= \left(\frac{5}{4}\right)^{-n} (E2^{X_1})(E2^{X_2}) \dots (E2^{X_n}) \quad (\text{because } X_1, X_2, \dots \text{ are independent}) \\ &= \left(\frac{5}{4}\right)^{-n} (E2^{X_1})^n \quad (\text{because } X_1, X_2, \dots \text{ have the same distribution}) \\ &= \left(\frac{5}{4}\right)^{-n} \left(2^1 \underbrace{P(X_1=1)}_{=\frac{1}{2}} + 2^{-1} \underbrace{P(X_1=-1)}_{=\frac{1}{2}}\right)^n \\ &= \left(\frac{5}{4}\right)^{-n} \left(\frac{5}{4}\right)^n = 1. \end{aligned}$$

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This contradicts (1).